

High dimensional statistics

Part 2: Shrinkage and Stein's phenomenon

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ENPC - February 2022

A small recap of Stein's phenomenon

Available on previous slides

- Gaussian sequence model: $\mathbf{y} \in \mathbb{R}^n$, $\mathbf{y} \sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2 I_n)$, we want to estimate $\boldsymbol{\mu}$.
- Linear estimators $\hat{\boldsymbol{\mu}}(\mathbf{y}) = A\mathbf{y}$, ($A = I_n$ is the ML solution)
- Error $\mathbb{E}_{\mathbf{y}}[\|\hat{\boldsymbol{\mu}}(\mathbf{y}) - \boldsymbol{\mu}\|^2]$
- Akaike information criterion (AIC) for linear estimators
- Taking $A = sI_n$ and minimizing with respect to s gives $\hat{\boldsymbol{\mu}}(\mathbf{y}) = \left(1 - \frac{n\sigma^2}{\|\mathbf{y}\|^2}\right)\mathbf{y}$

A small recap of Stein's phenomenon

On the board last time

- AIC for non-linear estimators = Stein's unbiased risk estimate
- Apply it to estimators of the form $\hat{\mu}(\mathbf{y}) = (1 - \frac{n\sigma^2}{\|\mathbf{y}\|^2})\mathbf{y}$
- Conclude that MLE is not admissible

Let's re-do these steps.

Gaussian integration by part

Lemma (Scalar to scalar function, standard Gaussian)

if $z \sim \mathcal{N}(0, 1)$ and $f : \mathcal{R} \rightarrow \mathcal{R}$ is differentiable such that

$\mathbb{E}[|f(z)|] < \infty$ and $\mathbb{E}[|f'(z)|] < \infty$ then $\mathbb{E}[zf(z)] = \mathbb{E}[f'(z)]$.

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Proof.

$$\mathbb{E}[zf(z)1(z > 0)] = \int_0^{\infty} z\phi(z)\left(\int_0^z f'(t)dt\right)dz \quad (1)$$

$$= \int_0^{\infty} f'(t)\left(\int_t^{\infty} z\phi(z)dz\right)dt \quad (2)$$

$$= \int_0^{\infty} f'(t)\phi(t)dt \quad (3)$$

$$= \mathbb{E}[1(z > 0)f'(z)] \quad (4)$$

Do the same with $z \leq 0$.

□

Gaussian integration by part

Lemma (Scalar to scalar function, scaled Gaussian)

if $z \sim \mathcal{N}(0, \sigma^2)$ and $f : \mathcal{R} \rightarrow \mathcal{R}$ is differentiable such that

$\mathbb{E}[|f(z)|] < \infty$ and $\mathbb{E}[|f'(z)|] < \infty$ then

$\mathbb{E}[zf(z)] = \sigma^2 \mathbb{E}[f'(z)]$.

Gaussian integration by part

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 if $z \sim \mathcal{N}(0, \sigma^2)$ and $f : \mathcal{R} \rightarrow \mathcal{R}$ is differentiable such that
 $\mathbb{E}[|f(z)|] < \infty$ and $\mathbb{E}[|f'(z)|] < \infty$ then
 $\mathbb{E}[zf(z)] = \sigma^2 \mathbb{E}[f'(z)]$.

Proof.

For $\sigma \neq 1$, take $g = z/\sigma$ and $\tilde{f}(u) = f(\sigma u)$

$$\mathbb{E}[g\tilde{f}(g)] = \mathbb{E}[\tilde{f}'(g)] \quad (5)$$

$$\iff \mathbb{E}[zf(z)] = \sigma^2 \mathbb{E}[f'(z)] \quad (6)$$

□

Gaussian integration by part

Lemma (vector to scalar function)

if $\mathbf{y} \sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2 I_n)$ and $g : \mathcal{R}^n \rightarrow \mathcal{R}$ is differentiable, integrable and so are the partial derivatives of g , then:

$$\mathbb{E}[(y_i - \mu_i)g(\mathbf{y})] = \sigma^2 \mathbb{E}\left[\frac{\partial g}{\partial y_i}(\mathbf{y})\right]$$

Gaussian integration by part

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Proof.

Take $z = (y_1 - \mu_1)$ and $f(z) = g(\mu_1 + z, y_2, \dots, y_n)$.

$$\mathbb{E}_z[zf(z)|y_{2:n}] = \sigma^2 \mathbb{E}_z[f'(z)|y_{2:n}] \quad (7)$$

$$\Rightarrow \mathbb{E}_{y_1}[(y_1 - \mu_1)g(\mathbf{y})|y_{2:n}] = \sigma^2 \mathbb{E}_{y_1}\left[\frac{\partial g}{\partial y_1}(\mathbf{y})|y_{2:n}\right] \quad (8)$$

$$\Rightarrow \mathbb{E}_{y_{2:n}}[\mathbb{E}_{y_1}[(y_1 - \mu_1)g(\mathbf{y})|y_{2:n}]] = \sigma^2 \mathbb{E}_{y_{2:n}}[\mathbb{E}_{y_1}\left[\frac{\partial g}{\partial y_1}(\mathbf{y})|y_{2:n}\right]]$$

□

Gaussian integration by part

Lemma (vector to vector function)

if $\mathbf{y} \sim \mathcal{N}(\boldsymbol{\mu}, \sigma^2 I_n)$ and $\mathbf{f} : \mathcal{R}^n \rightarrow \mathcal{R}^n$ is such that all coordinate functions are integrable with integrable partial derivatives then

$$\mathbb{E}[(\mathbf{y} - \boldsymbol{\mu})^\top \mathbf{f}(\mathbf{y})] = \sigma^2 \mathbb{E}[\sum_{i=1}^n \frac{\partial f_i}{\partial y_i}(\mathbf{y})]$$

Gaussian integration by part

Lemma (vector to vector function)

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$$\mathbb{E}[(\mathbf{y} - \boldsymbol{\mu})^\top \mathbf{f}(\mathbf{y})] = \sigma^2 \mathbb{E}[\sum_{i=1}^n \frac{\partial f_i}{\partial y_i}(\mathbf{y})]$$

Proof.

$$\mathbb{E}[(\mathbf{y} - \boldsymbol{\mu})^\top \mathbf{f}(\mathbf{y})] = \sum_{i=1}^n \mathbb{E}[(y_i - \mu_i) f_i(\mathbf{y})] \quad (9)$$

$$= \sum_{i=1}^n \sigma^2 \mathbb{E}\left[\frac{\partial f_i}{\partial y_i}(\mathbf{y})\right] \quad (10)$$

□

Stein's unbiased risk estimate

Theorem

Take $\hat{\boldsymbol{\mu}} = \mathbf{f}(\mathbf{y})$ and $\hat{r}_f = \|\mathbf{y} - \mathbf{f}(\mathbf{y})\|^2 + 2\sigma^2 \sum_{i=1}^n \frac{\partial f_i}{\partial y_i}(\mathbf{y}) - n\sigma^2$,
then $\mathbb{E}[\hat{r}_f] = \mathbb{E}[\|\boldsymbol{\mu} - \hat{\boldsymbol{\mu}}\|^2]$

Proof.

$$\mathbb{E}[\|\hat{\boldsymbol{\mu}} - \mathbf{y} + \mathbf{y} - \boldsymbol{\mu}\|^2] \tag{11}$$

$$= \mathbb{E}[\|\hat{\boldsymbol{\mu}} - \mathbf{y}\|^2 + \|\mathbf{y} - \boldsymbol{\mu}\|^2 + 2\langle \hat{\boldsymbol{\mu}} - \mathbf{y} | \mathbf{y} - \boldsymbol{\mu} \rangle] \tag{12}$$

$$= \mathbb{E}[\|\mathbf{f}(\mathbf{y}) - \mathbf{y}\|^2 + \|\boldsymbol{\epsilon}\|^2 + 2\langle \hat{\boldsymbol{\mu}} | \boldsymbol{\epsilon} \rangle - 2\langle \mathbf{y} | \boldsymbol{\epsilon} \rangle] \tag{13}$$

$$= \mathbb{E}[\|\mathbf{y} - \mathbf{f}(\mathbf{y})\|^2 + n\sigma^2 + 2\sigma^2 \sum_{i=1}^n \frac{\partial f_i}{\partial y_i}(\mathbf{y}) - 2n\sigma^2] \tag{14}$$

□

MLE is not admissible

Theorem

Take $\mathbf{f}(\mathbf{y}) = (1 - \frac{\sigma^2 q}{\|\mathbf{y}\|^2})\mathbf{y}$ Then

$$\mathbb{E}[\|\hat{\mu} - \mu\|^2] = \mathbb{E}[\hat{r}_f] = n\sigma^2 - \mathbb{E}[\frac{\sigma^4 q}{\|\mathbf{y}\|^2}(2n - q - 4)]$$

Proof.

- $\|\mathbf{y} - \mathbf{f}(\mathbf{y})\|^2 = \|\frac{\sigma^2 q}{\|\mathbf{y}\|^2}\mathbf{y}\|^2 = \frac{\sigma^4 q^2}{\|\mathbf{y}\|^2}$



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- $\|\mathbf{y} - \mathbf{f}(\mathbf{y})\|^2 = \|\frac{\sigma^2 q}{\|\mathbf{y}\|^2}\mathbf{y}\|^2 = \frac{\sigma^4 q^2}{\|\mathbf{y}\|^2}$
- $\frac{\partial f_i}{\partial y_i}(\mathbf{y}) = (1 - \frac{\sigma^2 q}{\|\mathbf{y}\|^2}) + \frac{2\sigma^2 q y_i^2}{\|\mathbf{y}\|^4}$



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- $\frac{\partial f_i}{\partial y_i}(\mathbf{y}) = (1 - \frac{\sigma^2 q}{\|\mathbf{y}\|^2}) + \frac{2\sigma^2 q y_i^2}{\|\mathbf{y}\|^4}$
- $2\sigma^2 \sum_{i=1}^n \frac{\partial f_i}{\partial y_i}(\mathbf{y}) = 2n\sigma^2(1 - \frac{\sigma^2 q}{\|\mathbf{y}\|^2}) + \frac{4\sigma^2 q}{\|\mathbf{y}\|^2} =$
 $2n\sigma^2 - \frac{\sigma^4 q}{\|\mathbf{y}\|^2}(2n - 4)$



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Take $\mathbf{f}(\mathbf{y}) = (1 - \frac{\sigma^2 q}{\|\mathbf{y}\|^2})\mathbf{y}$ Then

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Proof.

- $\|\mathbf{y} - \mathbf{f}(\mathbf{y})\|^2 = \|\frac{\sigma^2 q}{\|\mathbf{y}\|^2}\mathbf{y}\|^2 = \frac{\sigma^4 q^2}{\|\mathbf{y}\|^2}$
- $\frac{\partial f_i}{\partial y_i}(\mathbf{y}) = (1 - \frac{\sigma^2 q}{\|\mathbf{y}\|^2}) + \frac{2\sigma^2 q y_i^2}{\|\mathbf{y}\|^4}$
- $2\sigma^2 \sum_{i=1}^n \frac{\partial f_i}{\partial y_i}(\mathbf{y}) = 2n\sigma^2(1 - \frac{\sigma^2 q}{\|\mathbf{y}\|^2}) + \frac{4\sigma^2 q}{\|\mathbf{y}\|^2} = 2n\sigma^2 - \frac{\sigma^4 q}{\|\mathbf{y}\|^2}(2n - 4)$
- $\hat{r}_f = \|\mathbf{y} - \mathbf{f}(\mathbf{y})\|^2 + 2\sigma^2 \sum_{i=1}^n \frac{\partial f_i}{\partial y_i}(\mathbf{y}) - n\sigma^2 = n\sigma^2 - \frac{\sigma^4 q}{\|\mathbf{y}\|^2}(2n - q - 4)$

